

Effect of opportunistic scheduling on the quality of service perceived by the users in OFDMA cellular networks

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Abstract Our objective is to analyze the impact of *fading* and *opportunistic scheduling* on the quality of service perceived by the users in an orthogonal frequency-division multiple access cellular network. To this end, assuming Markovian arrivals and departures of customers that transmit some given data volumes, as well as some temporal channel variability (fading), we study the mean *throughput* that the network offers to users in the long run of the system. Explicit formulas are obtained in the case of *allocation policies*, which may or may not take advantage of the fading, called respectively *opportunistic* and *non-opportunistic*. The main practical results of the present work are the following: Firstly, we evaluate for the non-opportunistic allocation the *degradation due to fading* compared to additive white Gaussian noise (AWGN; that is, a decrease of at least 13% of the throughput). Secondly, we evaluate the *gain induced by the opportunistic* allocation. In particular, when the traffic demand per cell exceeds some value (about 2 Mbits/s in our numerical example), the gain induced by opportunism compen-

sates the degradation induced by fading compared to AWGN.

Keywords Radio communication · Opportunistic · Scheduling · Quality of service · Information theory · Markov processes

1 Introduction

Wireless cellular networks are in permanent and rapid evolution. They provide service for *constant bitrate* (CBR) calls (such as voice calls, video conferences, etc.) and carry *variable bitrate* (VBR) calls (such as mail, ftp), which accept fluctuations of the rates. Customers arrive to the network and require some service. Once they are admitted, they are allocated some *resource* (power, bandwidth, code) for the duration of their service, and then they depart from the network. Besides the *dynamics* due to user arrivals and departures, the channel conditions vary due to *fading*.

Growing traffic requires better planning and/or dimensioning of the networks. This task can be substantially simplified by analytical study of the dynamic network performance.

The objective of the present paper is to *account for the effect of fading in the dynamic performance evaluation* of the downlink of an orthogonal frequency-division multiple access (OFDMA) cellular network.

To this aim, the following two groups of elements are crucial: on one hand, the network geometry and the *resource allocation*, which can be chosen by the network designer, and on the other hand, *channel conditions* and the user traffic (intensity of arrivals and requested service) can only be predicted by the designer. In order

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to optimize the network performance, the resource allocation should be adapted to the channel and traffic conditions. Besides better coding (up to the *information theory's* limit) one is looking for allocations that are opportunistic, i.e., take advantage of the actual channel state and user position, preserving some fairness of the mean service rates offered to users. More precisely, the *opportunistic* allocation assumes that at each time, each bandwidth portion is allocated to the user with the most favorable fading state. Once the resource allocation is given, the quality of service (i.e., blocking probability for CBR and throughput for VBR) perceived by the users may be evaluated using *queueing theory*.

Individual elements of the above puzzle (i.e., information theory, resource allocation, and queueing theory) are often studied and optimized separately. The main contribution of this paper is a global approach that combines these elements to deduce the performance of the non-opportunistic and the opportunistic resource allocations. In doing so, we systematically use a separation of the times scales of different elements of the network dynamics, described in Section 2.1.

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The remaining part of this paper is organized as follows: In Section 1.1, we list these existing results in the literature that are used in our global approach. Our model of the OFDMA network is presented in Section 2. In Section 2.1, we describe the network dynamics and its decomposition into three time scales. They correspond, respectively, to information theory, resource allocation, and queueing theory which are studied in Sections 3, 4, and 5, respectively. The numerical results are described in Section 6. Finally, Section 8 describes some interesting perspectives.

1.1 Related work

A growing interest in OFDMA has resulted in many publications on performance evaluation of such networks. It is not in the scope of this paper to make a thorough review of this literature. For some references on this field, see for example those cited in [17]. Here we remark only those results that are directly related to our approach.

Our starting point is our work in [17] assuming AWGN channels between the users and the base stations. Coluccia et al. [11] study the optimality and fairness in resource allocation. As regard to the perfor-

mance of the opportunistic allocation policy over fading, we extend the approach of [6] and [19] to OFDMA networks. Moreover, we use the known results for the stability and stationary distribution of the multi-class extension of the processor sharing queue (see, e.g., [9, 21]).

The above results are used as building blocks to evaluate the dynamic performance of the non-opportunistic and the opportunistic resource allocations in OFDMA multi-cell networks. The present paper is an extended version of [5]. Besides more detailed proofs and analysis, the major extension is to show that the capacity lower bounds (which are the basis of our analytical approach) hold *true even if there are a few number of interferers* (Propositions 1 and 3 below). We carry this extension by building an argument based on the Jensen's inequality in place of the law of large numbers argument used in our previous work [5].

Some attempts to make other extensions of the present work are currently carried by other researchers; for example, Combes et al. [12] attempt to account for frequency selective fading channels.

2 Model assumptions

We will consider a wireless network composed of several base stations (BS). Each BS is equipped with a single antenna (no MIMO), and its total power is limited to some given maximal value. The same frequency spectrum is available to all BS (frequency reuse factor equal to one). Each BS allocates disjoint subcarriers to its users without macrodiversity (each user is served by exactly one BS). Thus, any given user receives only other-BS interference that is the sum of powers emitted by other BS on the subcarriers allocated to him by his BS.

When the radio signal propagates (from a BS to a given user), it is attenuated. We model the propagation loss as the product of three factors called distance loss, shadowing, and fading. The first factor is due to the *distance* between the transmitter and the receiver. The *shadowing* is due to the attenuation by the obstacles between the transmitter and the receiver. The *fading* is due to the reflection on the obstacles in the neighborhood of the receiver which generate multiple paths. We account only for the distance and the fading effects in the present study.

We assume that the bandwidth of each subcarrier is smaller than the *coherence frequency* of the channel, so that we can consider that the *fading* in each subcarrier is *flat*. That is, the output of the channel at a given time

depends on the input only at the same instant of time. We do not make any assumption on the *correlation* of the fading processes corresponding to different subcarriers (for a given user and a given BS). However, the fading processes for different users or base stations are assumed independent.

Time is divided into time slots of length smaller than the *coherence time* of the channel, so that, for a given subcarrier, the fading remains *constant during each time slot* and the fading process in different time slots may be assumed *ergodic* (such model for fading generalizes the so-called *quasi-static* model where the fading process at different time slots is assumed to be independent and identically distributed (i.i.d.)).

The fading in each time slot and each subcarrier is a *Rayleigh* distributed random variable. The code-word duration equals the time slot, which is assumed sufficiently large so that the *capacity* within each time slot may be defined in the *asymptotic* sense of the information theory.

Users perform *single user detection*; thus, the interference is considered as additive noise and added to the AWGN. The statistical properties of the interference are not known a priori since they depend on the codings of the other users. Moreover, the signals transmitted by different base stations are assumed independent.

For each subcarrier and each time slot, efficient (e.g., turbo) codes are used to obtain bitrates close to the capacity given by information theory (i.e., the maximal rate for which there exist coding schemes with error probability vanishing when the length of the code is sufficiently large). In particular, for the AWGN channel, the capacity equals

$$C = w \log_2(1 + \text{SNR}) \quad (1)$$

where w is the bandwidth and SNR designates the signal-to-noise power ratio.

2.1 Separation of time scales

We assume that there are a few time scales each associated to the evolution of some stochastic process. Roughly speaking, we assume that the duration of each time scale is sufficiently large so that the corresponding process converges to its stationary state and sufficiently small so that the processes of the larger time scales do not evolve within this duration.

We describe now the time scales from the fastest to the slowest one. The fastest time scale corresponds to *information theory*. The users' number and positions as well as powers and bandwidths allocated to them are

assumed fixed. Then the bitrate of each user cannot exceed the capacity given by information theory (i.e., the maximal rate for which there exist coding schemes with error probability vanishing when the length of the code is sufficiently large). This *information theory constraint* (i.e., existence of a coding scheme corresponding to the bitrate) together with the maximal power and bandwidth constraints constitute the *resource allocation constraints*.

The intermediate time scale corresponds to the *resource allocation* problem. The users' number and positions are assumed fixed. The network attempts to allocate powers, bandwidths, and bitrates (coding schemes) to all users respecting the resource allocation constraints.

This problem may be reformulated in terms of some condition on the bitrates. More precisely, by *feasibility condition* we mean a condition regarding bitrates of different users which guarantees the existence of powers and bandwidths satisfying the resource allocation constraints.

It is the role of the *load control* to guarantee this condition. The load control comprises the *admission control* (for CBR traffic) which checks at the arrival instant of each new user the *feasibility condition* and decides whether it can be admitted or not and the *congestion control* (for VBR traffic) which may modify the user bitrates to satisfy the *feasibility condition*. Once this is done, the network allocates powers and bandwidths to all users supporting the (feasible) bitrates.

The slowest time scale corresponds to *queueing theory*, where we consider the stochastic process of users' positions which is driven by *call arrivals and durations or data volumes*. This process is subject at each time the feasibility condition described above. The stationary state (or distribution) of this process permits to calculate the quality of service perceived by the users.

2.2 Power allocation

The network may operate a power adaptation. Such power adaptation may be either a *power control* to transmit just what is necessary to compensate the interference as described in [17] or a *water filling* attempting to maximize the cell capacity by allocating the powers among the users according to their fadings as described in [20] (Theorem 2, p. 119).

We shall assume that the signal transmitted by each BS has a *given power* constant over time (which equals to the maximal authorized value) and a *constant power spectral density* (i.e., the power is the same for all subcarriers). The performance obtained with this assumption gives a *lower bound* of the performance of a

network which operates a power adaptation. We shall carry our analysis by considering further lower bounds when necessary. The dimensioning based on such lower bounds is conservative.

2.3 Resource allocation

We always assume that each *user (receiver) knows* its own fading state, a situation called channel state information at receiver (CSIR) in literature.

We shall consider two resource allocation schemes:

1. *Non-opportunistic*: The BS does not take into account the fading when allocating the subcarriers to the users. This may be due to the fact that the BS (*transmitter*) *does not know* the fading states of its users.
2. *Opportunistic*: The BS takes into account the fading when allocating the subcarriers to the users. This requires that the BS (*transmitter*) *knows* the fading states of its users, a situation called channel state information at transmitter in literature. The gain obtained by the opportunistic allocation is called the *multiuser diversity* gain.

2.4 Possible model extensions

Even though our model seems to have many simplifying assumptions, we will show that it permits to capture and analyze the principal impacts of fading and multiuser diversity gain in cellular OFDMA networks. Moreover, the following remarks show how to extend the model.

In [15], it is observed that for OFDMA systems implementing a family of M -QAM modulations (as those described in [13]) with some BER target, the AWGN capacity formula 1 should be replaced by

$$C = w \log_2 \left(1 + \frac{\text{SNR}}{\Gamma} \right)$$

where $\Gamma = -\ln(5 \times \text{BER})/1.5$ (which is larger than 1 for $\text{BER} < e^{-1.5}/5 \simeq 0.0446$). Thus, accounting for real coding schemes may be tentatively taken into account in our approach by an appropriate modification of the AWGN capacity formula (for a similar question for high-speed downlink packet access (HSDPA) networks, it is observed in [16] (Fig. 11.1, p. 175) that HSDPA coding offers a capacity which is approximately the third of the AWGN capacity).

The assumption that the fading is constant over a time slot may be replaced by the less restrictive assumption that the fading is ergodic and perfectly known by the receiver (see [7]). In what follows, we analyze the three time scales of our model described in Section 2.1.

3 Information theory

We assume that the users do not move at the considered time scale. Consider a given user served by a given BS u . For each subcarrier and each time slot, say $[0, T]$, the channel output $Y(t)$ is related to the transmitted signal $X(t)$ by

$$Y(t) = S_u \times X(t) + Z(t) + I(t), \quad t \in [0, T]$$

where S_u represent the fading assumed constant, $Z(t)$ designate the noise assumed AWGN with power spectral density N_0 , and $I(t)$ is the interference (we skip for the moment the propagation loss induced by the distance).

Lets denote by p be the power in a given subcarrier, that is,

$$E[|X(t)|^2] = p$$

For each interfering BS v , let $X_v(t)$ be its transmitted signal and S_v be the fading assumed constant. Then the interference equals to

$$I(t) = \sum_{v \neq u} S_v X_v(t)$$

Since the signals transmitted by different base stations are assumed independent (and centered),

$$E[|I(t)|^2] = \sum_{v \neq u} |S_v|^2 E[|X_v(t)|^2] = p \sum_{v \neq u} |S_v|^2$$

Using [22] (Theorem 18), we may show that the worst noise process distribution (not necessarily white nor Gaussian) for capacity with given second moment is the AWGN. We deduce that the capacity C within the considered time slot is lower bounded by

$$C \geq w \log_2 \left(1 + \frac{p |S_u|^2}{w N_0 + p \sum_{v \neq u} |S_v|^2} \right)$$

where w is the bandwidth of a subcarrier.

Remark 1 An alternative way to show the above inequality is to use [1] (Lemma 1, p. 30) giving the capacity of the interference channel. It is frequent to see in literature the equality sign in place of the inequality in the above display, as for example [20] (p. 118). We do not see why the equality may be justified in this generality (see [2] (§4, p. 811) for a similar question).

We will now introduce propagation loss L_v induced by the distance between BS v and the given user. In

order to account for these losses, the above formula should be replaced by

$$C \geq w \log_2 \left(1 + \frac{p |S_u|^2 / L_u}{w N_0 + p \sum_{v \neq u} |S_v|^2 / L_v} \right) \quad (2)$$

3.1 Ergodic capacity

We will see soon that the expectation of the capacity with respect to the fading process (called *ergodic capacity*) plays a crucial role. We will not make any assumption of the distribution of the fading process except that

$$E[|S_u|^2] = 1$$

Proposition 1 *The ergodic capacity $E[C]$ is lower bounded by*

$$E[C] \geq w E[\log_2(1 + \beta |S_u|^2)]$$

where the expectation is with respect to the fading random variables and

$$\beta = \frac{p/L_u}{w N_0 + p \sum_{v \neq u} 1/L_v} \quad (3)$$

which may be viewed as the signal-to-interference and noise ratio (SINR) in our model.

Proof Note that the expectations in the present proof are with respect to the fading random variables S_u and $(S_v)_{v \neq u}$. Let $E[\cdot | S_u]$ designates the expectation conditionally to S_u . By the properties of the conditional expectation, we have

$$E[C] = E[E[C | S_u]]$$

Equation 2 implies

$$E[C | S_u] \geq w E \left[\log_2 \left(1 + \frac{\mathcal{P} |S_u|^2 / L_u}{N + \mathcal{P} \sum_{v \neq u} |S_v|^2 / L_v} \right) | S_u \right]$$

Since the function $x \mapsto \log(1 + 1/x)$ is convex on $\mathbb{R}_+^*$ (its second derivative is $\frac{2x+1}{x^2(x+1)^2}$), we deduce from Jensen's inequality that the right-hand side of the above equation is larger than $E[\log_2(1 + \beta |S_u|^2) | S_u]$ where β is given by Eq. 3. Thus,

$$\begin{aligned} E[C] &= E[E[C | S_u]] \\ &\geq w E[E[\log_2(1 + \beta |S_u|^2) | S_u]] \\ &= w E[\log_2(1 + \beta |S_u|^2)] \end{aligned}$$

□

4 Resource allocation

4.1 Non-opportunistic allocation

We consider a given subcarrier and multiple time slots. Recall that the fading for different time slots are independent and identically distributed. By the law of large number, the capacity averaged over a large number of time slots would approach the so-called *ergodic capacity* $E[C]$ where the expectation is with respect to the fading states. Proposition 1 gives a lower bound for $E[C]$.

Averaging over a large number of time slots corresponds to exploiting the so-called *time-diversity* which is suitable for the analysis of the performance of VBR traffic as observed in [8] (§I).

Remark 2 Note that the formula

$$E[\log_2(1 + \beta |S_u|^2)]$$

is similar to the capacity of a channel with ergodic fading in [7] (or i.i.d. fading in [14] (Eq. 2)) and CSIR (i.e., the receiver knows the fading). Nevertheless, the fading varies during a codeword there, whereas in the model considered in the present paper the fading is constant during a codeword.

Since $|S_u|$ is Rayleigh distributed, i.e., it has the probability density function

$$f_{|S_u|}(s) = 2se^{-s^2} 1_{s \geq 0}$$

then $H = |S_u|^2$ has exponential distribution

$$f_H(h) = e^{-h} 1_{h \geq 0}$$

In this case,

$$\begin{aligned} E[\log_2(1 + \beta |S_u|^2)] &= \int_0^\infty \log_2(1 + \beta h) e^{-h} dh \\ &= \frac{1}{\ln 2} e^{\beta^{-1}} \text{Ei}(1, \beta^{-1}) \end{aligned}$$

where

$$\text{Ei}(1, x) = \int_1^\infty \frac{e^{-xt}}{t} dt$$

is the *exponential-integral function*.

Remark 3 Frequency diversity. We consider now a given time slot and large number n of subcarriers. Assume that the fading variables for different subcarriers are i.i.d. (or more generally the fading process is ergodic with respect to the number n of subcarriers). Then

again, by the law of large number, the capacity of a large number n of subcarriers is lower bounded by

$$nwE[\log_2(1 + \beta |S_u|^2)]$$

Thus, the ergodic capacity is also appropriate for *CBR traffic* (where we cannot always count on time diversity) when the number of subcarriers allocated to each user is large enough.

4.1.1 Feasible bitrates

A base station u allocates to each user m in its cell (denoted with a slight abuse of notation by $m \in u$) a number of subcarriers of total bandwidth w_m (that is the bandwidth w of a given subcarrier multiplied by the number of subcarriers allocated to the user). The following proposition gives a condition on the users bitrates $(r_m)_{m \in u}$ assuring the feasibility of the bandwidths allocation.

Proposition 2 *Let u be a given base station, W be the total system bandwidth and*

$$\beta_m = \frac{p/L_{u,m}}{wN_0 + p \sum_{v \neq u} 1/L_{v,m}} \quad (4)$$

be the SINR for user $m \in u$, where N and $\mathcal{P}$ are, respectively, the noise and signal powers in a given subcarrier (of bandwidth w) and $L_{v,m}$ is the propagation loss between BS v and user m . If the users bitrates $(r_m)_{m \in u}$ satisfy the following condition:

$$\sum_{m \in u} \frac{r_m}{E[\log_2(1 + \beta_m |S_u|^2)]} \leq W \quad (5)$$

then the allocation of the bandwidths to the users is feasible.

Proof The feasibility of a bandwidth allocation may be formulated as follows: Do there exist bandwidths $(w_m)_{m \in u}$ whose sum does not exceed the total bandwidth W , that is,

$$\sum_{m \in u} w_m \leq W$$

We deduce from Proposition 1 that the following bitrates are feasible from the information theory point of view

$$r_m = w_m E[\log_2(1 + \beta_m |S_u|^2)], \quad m \in u$$

where the SINR β_m is given by Eq. 4. Since β_m is independent of the number of subcarriers allocated to the

user, we may easily express the bandwidth as function of the bitrate as follows:

$$w_m = \frac{r_m}{E[\log_2(1 + \beta_m |S_u|^2)]}$$

The condition 5 on the bitrates assures that the sum of the bandwidths given by the above equation does not exceed the total bandwidth W . $\square$

Note that Eq. 5 may be written as follows:

$$\sum_{m \in u} r_m \gamma_m \leq 1 \quad (6)$$

where

$$\gamma_m = [WE[\log_2(1 + \beta_m |S_u|^2)]]^{-1} \quad (7)$$

Each bitrate allocation satisfying Eq. 6 will be called *non-opportunistic allocation*. Note that the constraint 6 on the bitrates do not depend on the current values of the fading but only on their statistics.

4.2 Opportunistic allocation

We assume that the BS knows the fading states of all the users in its cell and that each user knows his own fading state.

4.2.1 Analysis for a given subcarrier

We account now for the fading variations at the resource allocation time scale. In what follows, by *opportunistic allocation* we mean that at a given time slot, each subcarrier is allocated to the user with the most favorable fading state; more precisely, each BS u allocates the subcarrier to the user m such that

$$|S_{u,m}| = \max_{n \in u} |S_{u,n}| \quad (8)$$

We aim now to calculate the ergodic throughput of each user per time slot under this policy.

If we allocate a given subcarrier to a user $m \in u$, then we deduce from Eq. 2 that, for a given time slot, we can find a coding scheme to support the following bitrate

$$R_m = w \log_2 \left(1 + \frac{p |S_{u,m}|^2 / L_{u,m}}{N + p \sum_{v \neq u} |S_{v,m}|^2 / L_{v,m}} \right), \quad m \in u$$

where the $S_{v,m}$ are the fading states between the user m and BS v for the considered subcarrier and time slot. Let us denote by $\bar{R}_m$ the bitrate effectively allocated by the opportunistic policy to a given user m averaged

over a large number of time slots. From the assumed temporal ergodicity, we deduce that

$$\bar{R}_m = E \left[R_m \times 1 \left\{ |S_{u,m}| = \max_{n \in u} |S_{u,n}| \right\} \right]$$

Proposition 3 Assume that for a given time slot the bit-rate of user m is given by

$$R_m = w \phi \left(\frac{p |S_{u,m}|^2 / L_{u,m}}{N + p \sum_{v \neq u} |S_{v,m}|^2 / L_{v,m}} \right), \quad m \in u \quad (9)$$

where ϕ is a function defined on $\mathbb{R}_+$ having a well-defined Laplace transform $\mathcal{L}_\phi(u) = \int_0^{+\infty} \phi(x) e^{-ux} dx$ and such that the function $x \mapsto \phi(1/x)$ is convex on $\mathbb{R}_+$. Then for the opportunistic allocation (Eq. 8), the bitrate effectively allocated to user m and averaged over a large number of time slots is lower bounded by

$$\bar{R}_m \geq R'_m \quad (10)$$

$$\begin{aligned} &:= w E \left[\phi \left(\beta_m |S_{u,m}|^2 \right) \times 1 \left\{ |S_{u,m}| = \max_{n \in u} |S_{u,n}| \right\} \right] \\ &= w \int_0^{+\infty} \phi(\beta_m x) e^{-x} (1 - e^{-x})^{M-1} dx \end{aligned} \quad (11)$$

$$= \frac{w}{\beta_m} \sum_{k=0}^{M-1} \binom{M-1}{k} (-1)^k \mathcal{L}_\phi \left(\frac{k+1}{\beta_m} \right) \quad (12)$$

where β_m is given by Eq. 4 and M is the number of the users in the cell.

Proof Since the function $x \mapsto \log(1 + 1/x)$ is convex, the inequality 10 follows from Jensen's inequality. Since the $|S_{u,n}|^2$ are exponentially distributed, we get

$$\begin{aligned} R'_m &= w \int_{\mathbb{R}_+^M} \phi(\beta_m x_m) 1 \left\{ x_m = \max_{n \in u} x_n \right\} \prod_{n \in u} e^{-x_n} dx_n \\ &= w \int_{\mathbb{R}_+} \phi(\beta_m x) e^{-x} \int_{\mathbb{R}_+^{M-1}} \prod_{n \neq m} 1 \{x_n \leq x\} e^{-x_n} dx_n dx \\ &= w \int_{\mathbb{R}_+} \phi(\beta_m x) e^{-x} \prod_{n \neq m} \int_{\mathbb{R}_+} 1 \{x_n \leq x\} e^{-x_n} dx_n dx \\ &= w \int_{\mathbb{R}_+} \phi(\beta_m x) e^{-x} (1 - e^{-x})^{M-1} dx \\ &= w \int_0^{+\infty} \phi(\beta_m x) e^{-x} (1 - e^{-x})^{M-1} dx \end{aligned}$$

Using the binomial formula

$$(1 - e^{-x})^{M-1} = \sum_{k=0}^{M-1} \binom{M-1}{k} (-1)^k e^{-kx}$$

we get

$$\begin{aligned} R'_m &= w \sum_{k=0}^{M-1} \binom{M-1}{k} (-1)^k \int_0^{+\infty} \phi(\beta_m x) e^{-(k+1)x} dx \\ &= w \sum_{k=0}^{M-1} \binom{M-1}{k} (-1)^k \int_0^{+\infty} \phi(y) e^{-\frac{k+1}{\beta_m} y} \frac{dy}{\beta_m} \\ &= \frac{w}{\beta_m} \sum_{k=0}^{M-1} \binom{M-1}{k} (-1)^k \mathcal{L}_\phi \left(\frac{k+1}{\beta_m} \right) \end{aligned}$$

□

We prove now that the Shannon function $\log_2(1+x)$ satisfies the conditions of the above proposition.

Lemma 1 The function $\phi(x) = \log_2(1+x)$ has a well defined Laplace transform given by

$$\mathcal{L}_\phi(u) = \frac{e^u}{u \ln 2} \text{Ei}(1, u) \quad (13)$$

and the function $x \mapsto \phi(1/x)$ is convex on $\mathbb{R}_+$.

Proof The Laplace transform of $\phi(x) = \log_2(1+x)$ is

$$\begin{aligned} \mathcal{L}_\phi(u) &= \int_0^{+\infty} \log_2(1+x) e^{-ux} dx \\ &= \frac{1}{\ln 2} \int_0^{+\infty} \ln(1+x) e^{-ux} dx \\ &= \frac{1}{\ln 2} \int_0^{+\infty} \frac{e^{-ux}}{u(1+x)} dx \\ &= \frac{e^u}{u \ln 2} \int_1^\infty \frac{e^{-ut}}{t} dt = \frac{e^u}{u \ln 2} \text{Ei}(1, u) \end{aligned}$$

where for the third line we use an integration by parts. The convexity of the function $x \mapsto \phi(1/x)$ on $\mathbb{R}_+$ was already shown in the proof of Proposition 1.

4.2.2 Symmetric case

The expression of R'_m (defined by Eq. 10) in the case when ϕ is linear; that is, $\phi(x) = x/\ln 2$ was already given by Borst [6] (§II) who showed that

$$R'_m = \left(\frac{1}{M} \sum_{k=1}^M \frac{1}{k} \right) \frac{w \beta_m}{\ln 2} \quad (14)$$

This case is called *symmetric* in [6] (§II) since the random variables $\phi(\beta_m |S_{u,m}|^2)$ normalized by their expectations with respect to the fading states are identically distributed.

On the other hand, $\mathcal{L}_\phi(u) = \frac{1}{u^2 \ln 2}$; thus, we deduce from Eq. 12 that

$$R'_m = \left(\sum_{k=0}^{M-1} \binom{M-1}{k} (-1)^k \frac{1}{(k+1)^2} \right) \frac{w\beta_m}{\ln 2}$$

which is coherent with Eq. 14 since it is straightforward to prove that

$$\sum_{k=0}^{M-1} \binom{M-1}{k} (-1)^k \frac{1}{(k+1)^2} = \frac{1}{M} \sum_{k=1}^M \frac{1}{k}$$

(indeed, letting $f(x) = \sum_{k=0}^{M-1} \binom{M-1}{k} (-1)^k \frac{e^{-(k+1)x}}{(k+1)^2}$ and observing that $f''(x) = e^{-x} (1 - e^{-x})^{M-1}$, we get $f(x) = \frac{1}{M} \sum_{k=1}^M \frac{1 - (1 - e^{-x})^k}{k}$ which applied to $x = 0$ gives the desired result).

Thus, the expression of R'_m in our Proposition 3 may be seen as an extension of the result in [6] (§II) to the non-symmetric case, since it holds true even when the random variables $\phi(\beta_m | S_{u,m}|^2)$ normalized by their expectations are not necessarily identically distributed.

Note finally that Eq. 14 may be written as follows:

$$R'_m = \frac{1}{h(M)} \frac{w}{W} \frac{1}{\gamma'_m} \quad (15)$$

where $h(M) = M / \sum_{k=1}^M 1/k$ and $\gamma'_m = \frac{\ln 2}{W\beta_m}$. Thus, in the symmetric case, R'_m is decomposed into some function the number M of the users multiplied by some function of the position of user m . This particular form allows an explicit calculus of the performance at the queueing theory time scale as will be shown in Proposition 5 below.

4.2.3 Feasible bitrates

The bitrate for user m over all the bandwidth, say r'_m , is related to the bitrate R'_m per subcarrier by

$$r'_m = \frac{W}{w} R'_m \quad (16)$$

where R'_m is given by Eq. 11.

Example 1 In the particular case when R'_m is given by Eq. 15, then

$$r'_m = \frac{1}{h(M)} \frac{1}{\gamma'_m} \quad (17)$$

5 Queueing theory

5.1 Traffic dynamics

Denote the geographic region covered by the a cell by $\mathbb{D}$ that is assumed to be a bounded subset of the plane $\mathbb{R}^2$. Consider only VBR calls whose inter-arrival times to $\mathbb{D}$ are i.i.d. exponential random variables with rate λ (mean $1/\lambda$). The position of each arrival is picked at random in $\mathbb{D}$ according to the uniform distribution. We assume that users do not move during their calls. Each call requires to transmit a given volume of data (amount of bits that has to be sent or received), which is modeled by an independent of everything else, exponential random variable with parameter μ . The quantity $\rho = \lambda/\mu$ is called the *traffic demand* (expressed in Mbps¹) per cell. Users are served by the BS according to some bitrate allocation policy.

The set of positions of all users served at a given time is called *configuration of users*. Let $\mathbb{M}$ be the set of all possible configurations (this can be formalized, e.g., on the basis of the theory of point processes). We denote by $\{N_t\}_{t \geq 0}$ the process describing the evolution in time of the user configurations in $\mathbb{D}$ (due to arrivals and departures). It takes its values in $\mathbb{M}$. If the process $\{N_t\}_{t \geq 0}$ is not ergodic, then the mean number of users in the system grows unboundedly in the long run of the system. This situation has to be avoided, in which case we say that the system is *stable* (or equivalently ergodic).

One distinguish two milestones of the analytical *evaluation of the network performance*: identification of its stability region and the evaluation of the steady-state characteristics (e.g., the mean throughput and delay).

5.2 Non-opportunistic allocation

Define the *critical traffic demand* ρ_c by

$$\rho_c = \frac{1}{\bar{\gamma}}$$

where $\bar{\gamma}$ is the expectation of γ_m (which is given by Eq. 7) with respect to the position of the user m uniformly distributed in the cell.

Proposition 4 Assume any non-opportunistic work-conserving bitrate allocation satisfying condition 6. If the traffic demand per cell

$$\rho < \rho_c \quad (18)$$

¹The abbreviation Mbps designates “kilobit per second”.

then the system is stable. Consider the particular allocation

$$r_m = \frac{1}{M\gamma_m} \quad (19)$$

Then, at the steady state, the expected number of users in cell u , denoted $\bar{N}$, equals

$$\bar{N} = \frac{\rho}{\rho_c - \rho} \quad (20)$$

the mean delay per call, denoted $\bar{T}$, is

$$\bar{T} = \frac{1}{\mu\bar{r}} = \frac{1}{\mu(\rho_c - \rho)} \quad (21)$$

and the mean throughput, denoted $\bar{r}$, equals

$$\bar{r} = \rho_c - \rho \quad (22)$$

Proof A bitrate allocation is said to be *work conserving* when $\sum_{m \in u} r_m \gamma(m) = 1$ as long as there is at least a user in the cell. The stability condition 18 and the expression 20 of the users number at the steady state follow from known results for multi-class processor sharing queues [10, 18]. Applying Little's formula [3], we get the expression 21 for the delay. Using the fact that the throughput $\bar{r}$ is the ratio of the data volume average, which equals $1/\mu$, and the delay $\bar{T}$ proves Eq. 22. $\square$

5.3 Opportunistic allocation

The following proposition gives the performance of the opportunistic allocation (Eq. 17):

Proposition 5 Consider the opportunistic bitrate allocation (Eq. 17) and denote by $\bar{\gamma}'$ the expectation of γ'_m with respect to the user's location m uniformly distributed in the cell. If the traffic demand per cell

$$\rho < \frac{1}{\bar{\gamma}'} \lim_{n \rightarrow \infty} \frac{n}{h(n)}$$

then the system is stable. Moreover, at the steady state, the mean throughput per user, denoted $\bar{r}'$, equals

$$\bar{r}' = \frac{1}{\bar{\gamma}'} \frac{1}{\mathcal{H}(\bar{\gamma}'\rho)}$$

where the function $\mathcal{H}(s)$ is defined for $s > 0$ by

$$\mathcal{H}(s) = \frac{E[H(X+1)]}{E[H(X)]}, \quad H(M) = \begin{cases} \prod_{k=1}^M h(k) & \text{if } M \geq 1 \\ 1 & \text{if } M = 0 \end{cases}$$

where X is a Poisson random variable with parameter s . Moreover, the mean delay is

$$\bar{T} = \frac{1}{\mu\bar{r}'} = \frac{\bar{\gamma}'\mathcal{H}(\bar{\gamma}'\rho)}{\mu}$$

and the mean number of users at the steady-state equals

$$\bar{\gamma}'\rho\mathcal{H}(\bar{\gamma}'\rho)$$

Proof See [6] (Proposition 3.1) and [16] (§11.1). $\square$

5.3.1 Approximation

Note that Proposition 5 assumes a particular case when the decomposition (Eq. 17) holds true as in the symmetric case described in Section 4.2.2.

Unfortunately, for the more general allocation (Eq. 16), we do not know a closed form for the performance, especially because the geometric parameters β_m depend on the location of the user m . The idea is then to replace these parameters by a constant β_0 calculated as follows.

β_0 is determined in such a way that the average over the cell of the factors γ_m given by Eq. 7 should remain unchanged when the β_m are replaced by β_0 (by Proposition 4, this is equivalent to say that the performance of the non-opportunistic allocation (Eq. 19) remains unchanged when the β_m are replaced by β_0).

We will approximate the performance of our opportunistic scheme (Eq. 16) by replacing β_m with β_0 and applying Proposition 5.

6 Numerical results

6.1 Model specification

In order to obtain numerical values, we consider the most popular *hexagonal network model*, where the base stations are placed on a regular infinite hexagonal grid. Let R be the radius of the disk whose area is equal to that of the hexagonal cell served by each base station and call R the *cell radius*. We take $R = 0.525$ km. We assume a distance loss $L(r) = (Kr)^\eta$, with $\eta = 3.38$, $K = 8,667 \text{ km}^{-1}$. This means that the distance loss between BS u and user m is equal to $L(|x_u, x_m|)$ where x_u, x_m denote, respectively, the geometric location of u and m and $|\cdot|$ is the Euclidean distance.

We consider VBR traffic with traffic demand assumed (spatially) uniform over the cell. We consider a traffic demand per cell ρ varying from 0 to 20 Mbps.

The BS maximal total power is $\bar{P} = 52$ dBm (accounting for antenna gains and losses), the total

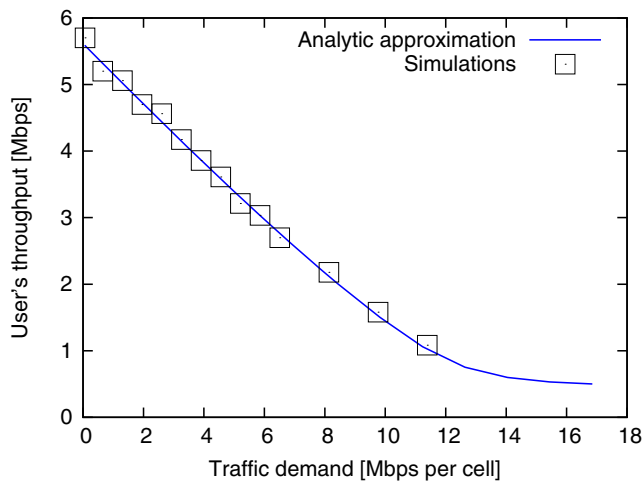


Fig. 1 Validation of the approximation for the opportunistic allocation

bandwidth is $W = 5$ MHz, and the ambient noise power $WN_0 = -103$ dBm.

6.2 Results

Figure 1 shows the throughput per user as function of the traffic demand per cell for the opportunistic allocation. We show the result of the simulations and of the approximation described in Section 5.3.1. The two curves are close which validates the considered approximation.

Figure 2 shows the throughput per user as function of the traffic demand per cell for the AWGN, non-opportunistic and opportunistic allocation. Firstly, we

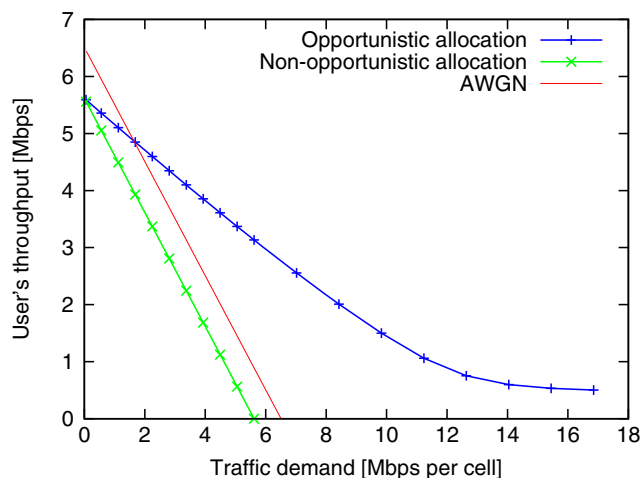


Fig. 2 Comparison of the performance of the different allocations

observe that for the non-opportunistic allocation, the fading *decreases* the throughput per user of at least 13% compared to AWGN. Note moreover that the non-opportunistic allocation gives, like AWGN, the throughput that is a linear function of the demanded traffic in the bounded stability region. Secondly, the *gain* induced by opportunism increases with the traffic. When the traffic demand is very small, the gain induced by the opportunistic allocation is negligible. Nevertheless, when the traffic demand per cell exceeds some value (about 2 Mbps in our example), the gain induced by opportunism compensates the degradation induced by fading compared to AWGN.

7 Conclusion

We have given analytic expressions for the mean throughput and delay in OFDMA cellular networks with VBR traffic. These formulas take into account in a simple but not simplistic way all important elements of the network performance: the network geometry, the traffic dynamics, fading, and bitrate allocation policy. They allow, in particular, to compare in a systematic way the effect on performance of fading for either non-opportunistic or opportunistic allocation policies.

8 Perspectives

Even though we assumed Rayleigh distribution for the fading, the approach may be extended to other distributions such as Rice, Nakagami, etc.

It may be useful to study the effect of grouping the subcarriers into the so-called *chunks* (i.e., a group of subcarriers and a number of time slots; see [20] (§6.2)).

As observed in [20] (§6.2), the non-opportunistic allocation scheme is suitable for fast varying fading, due to *high mobility* of users for example, which makes difficult to track of the rapidly varying channel. On the other hand, the opportunistic allocation scheme is suitable for slow varying fading as in the case of *fixed (or pedestrian)* users for example. Thus, from the point of view of the present paper, mobility decreases performance. It is interesting to couple this with the account of the speed effect at the queueing theory time scale in [4] which have shown that mobility ameliorates performance.

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